

# Long-term Air Pollution Exposure and Incident Dementia: A Systematic Review and Meta-Analysis

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### **Panel: Research in context**

#### **Evidence before this study**

Before undertaking this study, we reviewed the existing literature, including systematic reviews and meta-analyses, on the associations between various air pollutants and dementia incidence. We found key gaps, such as limited evidence for NO<sub>2</sub>, NO<sub>x</sub>, and BC/PM<sub>2.5</sub> Absorbance, a lack of certainty of evidence assessment, and few subgroup or sensitivity analyses to explore the sources of large observed heterogeneity.

Additionally, we assessed that the number of relevant primary studies had grown tremendously since the previous systematic reviews' publications. We then developed and published a protocol to guide our work.

This systematic review and meta-analysis searched Medline, Embase, Cochrane Library, Cinahl, Global Health, PsycINFO, Scopus, and Web of Science, covering studies through October 2023. Eligible studies were observational epidemiological studies that included participants > 18, examined long-term outdoor air pollution exposure (≥1 year), and quantitatively investigated exposure to a clinical dementia diagnosis in adults without dementia at baseline. Excluded studies exclusively reported on cognitive impairment, surrogate dementia markers (e.g., imaging, and biochemical markers), or lacked quantitative exposure or exposure-response estimates. We assessed risk of bias and certainty of evidence using standard tools from the OHAT Handbook.

Out of 15,619 screened records, 51 studies met the inclusion criteria, 40 were eligible for the meta-analysis, and 34 were included (population represented: 29,364,106). The meta-analyses revealed significant positive associations between PM<sub>2.5</sub> (Hazard Ratio (HR) per 5 µg/m<sup>3</sup>: 1.08 (95%CI: 1.02-1.14), BC/PM<sub>2.5</sub> absorbance (per 1 µg/m<sup>3</sup>: 1.13 (95%CI: 1.01-1.27) and NO<sub>2</sub> (per 10 µg/m<sup>3</sup>: 1.03 (95%CI: 1.01-1.05) and dementia incidence. Our certainty of evidence assessment indicated that we had a moderate level of certainty. Our risk of bias assessment of each included study was generally low, though a few studies had high bias in specific categories.

#### **Added value of this study**

This study offers the most comprehensive exposure-response estimates to date for use in burden of disease, health impact assessments and air quality regulation. It is the largest meta-analysis to explore the associations between air pollution and dementia incidence, covering pollutants not previously analysed. Subgroup analyses by dementia subtype, study location, outcome ascertainment and exposure assessment methods provide new insights into why evidence strength may vary across groups.

#### **Implications of all the available evidence**

Our findings reinforce and extend the evidence that long-term exposure to outdoor air pollution is a risk factor for the onset of dementia. Reducing pollution exposure would lower dementia rates, and stricter air quality standards are likely to yield significant health, social, and economic benefits.

# Abstract

A rapidly-evolving evidence-base suggests that exposure to air pollution increases dementia's incidence. We conducted a systematic review and meta-analysis according to a published protocol. Prespecified inclusion criteria required primary observational studies of individuals over 18 years and a quantitative report on the association between long-term outdoor air pollution and subsequent dementia incidence. Risk of bias, publication bias, and confidence in the evidence were assessed. Fifty-one papers met the inclusion criteria and 40 were eligible for some meta-analysis, with a population of 29,364,106 participants,<sup>1</sup> mostly from high-income countries. Of those 40 papers, 34 papers with continuous effect estimates whose pollutant-exposure met the threshold of three or more papers were included in the meta-analysis. Meta-analyses revealed significant positive associations between PM<sub>2.5</sub>, NO<sub>2</sub>, and BC/PM<sub>2.5</sub> Absorbance and dementia incidence. There was considerable heterogeneity between studies in the pooled analyses for most exposure-outcome associations. Sensitivity analyses supported findings. Subgroup analyses suggested a larger association between PM<sub>2.5</sub> and NO<sub>2</sub> with vascular dementia than Alzheimer's disease. The overall Certainty of Evidence was moderate. This paper adds to the body of evidence that outdoor air pollutants are risk factors for dementia.

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<sup>1</sup>This is a sum of the populations, rather than samples, represented by the studies included in the meta-analysis.

# 1 **Body**

## 2 **Introduction**

3 Dementia refers to the acquired and often progressive loss of function across multiple cognitive domains  
4 such that it impairs an individual's personal, social, and occupational function (1). Dementia is a clinical  
5 syndrome which can be caused by multiple disease processes. Estimated to affect over 57.4 million people,  
6 Alzheimer's disease (AD) and other dementias ranked as the eighth leading cause of death in 2021 globally  
7 (2, 3). Its global burden is projected to increase to 152.8 million cases by 2050 (3). The impacts on both  
8 society and the individual are significant, and, while there are some indications that its prevalence is  
9 decreasing in some higher income countries, signalling the potential to modify population risk, there are no  
10 signs that the incidence of dementia decreasing outside of Europe and North America (4). Therefore, there  
11 remains an urgent need to investigate and address causes of this syndrome.

12  
13 The most recent Lancet Commission on Dementia Prevention, Intervention, and Care (2024) identified air  
14 pollution as one of 14 modifiable risk factors for dementia, based on published evidence (5). Air pollution  
15 is a relatively recently recognised risk factor, supported by primary studies reporting an association between  
16 dementia and exposure to a number of air pollutants, including particulate matter (PM) of different sizes  
17 (PM<sub>2.5</sub>, PM<sub>2.5-10</sub>, PM<sub>10</sub>), Black Carbon (BC) or PM<sub>2.5</sub> Absorbance, and gases (ground-level Ozone (O<sub>3</sub>)),  
18 Nitrogen Dioxide (NO<sub>2</sub>), Carbon Monoxide (CO), Sulphur Oxides (SO<sub>x</sub>) and Nitrogen Oxides (NO<sub>x</sub>). The  
19 direction of association and strength of evidence, however, have varied by publication and pollutant, raising  
20 questions about confidence in the evidence base and reasons underlying observed heterogeneity (6-10).

21 Considering the then-emerging nature of this investigation and mixed findings of primary studies, a series  
22 of systematic reviews and meta-analyses were published from 2015 to synthesise the rapidly growing body  
23 of evidence. These works have concluded that PM<sub>2.5</sub> exposure is associated with an increased risk of  
24 developing dementia but yielded less clear findings on the associations with PM<sub>2.5-10</sub>, PM<sub>10</sub>, PM<sub>2.5</sub> Absorbance,

25 BC, O<sub>3</sub>, NO<sub>2</sub>, CO, SO<sub>x</sub>, and NO<sub>x</sub> and the differential effects on study-diagnosed dementia subtype (11-20).  
26 Importantly, many of these works were either systematic reviews without meta-analyses, included few  
27 primary studies in their meta-analysis, had inadequate definitions of case ascertainment, focused  
28 exclusively on the relationship between dementia and one (e.g., PM<sub>2.5</sub>), or a limited number of air pollutants,  
29 and differed in the strength, significance and direction of association (12, 13, 15-17, 20). Some of these  
30 studies focus on particular study-diagnosed subtypes (labelled as AD and vascular dementia (VaD), for  
31 example), which are subject to considerable uncertainty when tested against pathological findings after  
32 death, including the fact that most dementia in the ninth decade and beyond is mixed in nature, and such  
33 pathologies are found in most people who die without having developed dementia (21, 22).

34  
35 The latest systematic review and meta-analysis were published by Wilker et al. and included 51 studies  
36 published up to July 2022, only 16 of which were included in the meta-analysis (19). The authors concluded  
37 that there was evidence to support an association between dementia and PM<sub>2.5</sub> (14 studies), and more limited  
38 support for an association with NO<sub>2</sub> (9 studies) and NO<sub>x</sub> (5 studies). Additionally, they did not find a clear  
39 association with O<sub>3</sub> (4 studies). They assessed most studies to be at high risk of bias. This study provided  
40 important estimates but was limited in the number of pollutants included in its meta-analysis with limited  
41 data for NO<sub>2</sub>, and no data for PM<sub>10</sub>, or BC/PM<sub>2.5</sub> absorbance. These pollutants are included in our meta-analysis.  
42 Secondly, Wilker et al. did not perform a subgroup analysis by dementia study-diagnosed subtype. Finally,  
43 it lacked an overall certainty of evidence assessment and a publication bias assessment, which the authors  
44 noted as a limitation (19).

45  
46 Since 2022, a surge in publications has provided an opportunity for a new, more comprehensive systematic  
47 review and meta-analysis. The primary aim of this study is to deepen the understanding of the relationship  
48 between air pollution and dementia incidence, while bolstering the evidence base for strategies to prevent  
49 or delay its onset. New studies include additional pollutants like NO<sub>2</sub>, NO<sub>x</sub>, PM<sub>10</sub>, and BC/PM<sub>2.5</sub> absorbance,  
50 offering a larger data pool for analysis, including subgroup effects, dementia subtype variations, and

51 potential clarification of earlier inconsistencies and heterogeneities. This expanded data also facilitates  
52 assessments of study quality and evidence certainty, and highlights areas for methodological improvement  
53 in future research. Ultimately, this study provides updated estimates for future burden of disease and health  
54 impact assessments and inform discussions on the mechanisms by which air pollution contributes to  
55 dementia, as well as policy on air quality limits and targets.

56

## 57 **Methods**

58 A protocol for this systematic review was published in December 2022 (23), and registered with the  
59 International Prospective Register of Systematic Reviews (PROSPERO) (CRD42023414413). The  
60 deviations from this protocol are listed in Table S1 (Supplementary Material pg. 1) and are mostly minor  
61 and more inclusive in nature.

62

## 63 **Literature Search**

### 64 *Search Strategy*

65 Medline (via Ovid), Embase (via Ovid), Cochrane Library, Cinahl (via Ebscohost), Global Health (via  
66 Ebscohost), PsycINFO (via Ebscohost), Scopus, and Web of Science (Core Collection) were searched twice  
67 (once from inception to October 2022 and again from October 2022 to October 2023) using peer-reviewed  
68 and piloted search terms (S1; Supplementary Material pg. 10) developed through a review of existing  
69 literature, a selection of indicator papers, and pilot searches, according to the published protocol (23).

70

### 71 *Eligibility Criteria*

72 Studies were eligible if they were based on the results of a primary case-control, cohort, cross-sectional, or  
73 ecological studies of individuals over 18 years of age, investigated exposure to outdoor air pollution of a  
74 year or more (long-term), and quantitatively reported the association between single air pollution exposures  
75 and a subsequent physician diagnosis of dementia in adults free of disease at baseline. To be included, cross

76 sectional studies must have a design adequate to answer the research question ‘is there an association  
77 between preceding exposure to air pollution and subsequent risk of the selected health outcome reported as  
78 incidence or life-time prevalence in healthy/disease-free individuals at baseline?’. In cross-sectional studies,  
79 only life-time prevalence was acceptable in previously healthy/disease-free individuals at baseline. We  
80 excluded papers that exclusively used outcomes such as cognitive impairment or self-diagnosis of dementia,  
81 and papers that did not provide a quantitative exposure estimate or used a proximity model for exposure  
82 assessment. Complete inclusion and exclusion criteria are in Table S2 (Supplementary Material pg. 2) and  
83 the published protocol (23).

84

#### 85 *Selection Process*

86 Deduplication of the search results was completed in EndNote then Zotero and then Rayyan. Remaining  
87 papers were title and abstract screened by two independent reviewers. Papers included by both independent  
88 reviewers, or included by one and excluded by the other, underwent full-text screening by two reviewers;  
89 any remaining conflict was resolved by a third reviewer. The list of studies excluded at both the title and  
90 abstract (Table S7) and full text screening stages (Table S8), alongside the reason for exclusion are included  
91 in the supplementary material. The original search included multiple sclerosis, Parkinson’s disease, and  
92 amyotrophic lateral sclerosis, the topic of another paper, alongside dementia. Following title and abstract  
93 screening, papers were divided by outcome for separate data extraction and analysis.

94

#### 95 *Data Extraction*

96 Using a standardised and piloted form through multiple iterations (Table S9), data was extracted  
97 independently by two reviewers. Duplicate sheets were compared, and discrepancies were resolved through  
98 escalation to and discussion with a third reviewer. Unadjusted and adjusted measures of association (effect  
99 estimates) were recorded with their reported 95% confidence intervals, unit of exposure ( $\mu\text{g}/\text{m}^3$ , ppb, etc),  
100 scaling factor (e.g.,  $1 \mu\text{g}/\text{m}^3$ ,  $5 \mu\text{g}/\text{m}^3$ ,  $10 \mu\text{g}/\text{m}^3$ ), and covariates adjustment. Additional data on study type,

101 location, year, length of follow-up, population, exposure assessment, and funding were also recorded. The  
102 complete list of extracted data points is included in S2 (Supplementary Material pg. 15) and includes 41  
103 items per paper. If one paper reported more than one effect estimate per pollutant, the estimate which met  
104 pre-specified criteria in the published protocol (Table S3; Supplementary Material pg. 2) was extracted  
105 (23). For example, if multiple adjusted effect estimates were reported for an individual pollutant, we  
106 extracted the effect estimate with the most restrictive analysis model (i.e., which adjusted for the greatest  
107 number of factors), if that adjustment was appropriate. We considered an adjustment to be inappropriate if  
108 it included intermediate conditions and/or effect modifiers such as stroke, heart failure, ischemic heart  
109 disease, diabetes, and depression. If the only available adjusted effect estimates included such  
110 comorbidities, these effect estimates were used, per the published protocol (23). A sensitivity analysis was,  
111 however, later run excluding effects estimates which adjusted for comorbidities from the meta-analysis.

112

113 If information could not be determined for a paper, we attempted to contact the authors to clarify, and if  
114 no response was received after three attempts, we coded it as 'Not Available'.

115

### 116 *Risk of Bias Assessment*

117 Using the Office of Health Assessment and Translation (OHAT) tool, two independent reviewers assessed  
118 Risk of Bias (RoB) for each included study (internal validity), with modifications implemented after  
119 piloting the tool at the outset (24). While the OHAT tool includes 11 questions, 4 of these questions were  
120 excluded, because they are related to Experimental Animal and Human Controlled Trials, making them  
121 inapplicable. For each of the 7 remaining questions, one of the following four answers was selected:  
122 definitely low risk of bias, probably low risk of bias, probably high risk of bias, definitely high risk of bias.  
123 Discrepancies between reviewers were discussed and a consensus was reached and, therefore, no conflicts  
124 or discrepancies needed to be escalated to the third senior reviewer, per the protocol. The full RoB  
125 assessment is shown in Table S4 (Supplementary Material pg.3). *Certainty of Evidence Assessment*

126 Two reviewers rated the confidence in evidence bodies using the OHAT approach for systematic review  
127 and evidence integration for literature-based health assessments (24). The studies were grouped together  
128 and given an initial confidence rating of “high confidence”, “moderate confidence”, “low confidence”, or  
129 “very low confidence” based on study design features which indicate the ability of a study to assess  
130 causality. Because observational studies do not have “controlled exposure”, one of the four key features in  
131 rating certainty of evidence, the highest rating the body of evidence could attain was “moderate confidence”  
132 provided they had the other three categories: “exposure prior to outcome”, “individual outcome data”, and  
133 “comparison group used”. Since our eligibility criteria selected for certain study design features, the body  
134 of evidence had an initial confidence rating of moderate, as all included studies had the other three  
135 categories. The confidence rating of the body of studies was upgraded or downgraded based on the presence  
136 or absence of various factors. The presence of factors which led to a downgrade in the confidence rating  
137 were risk of bias, unexplained consistency, indirectness, imprecision, and publication bias. The presence of  
138 factors which led to an upgrade in the confidence rating were a large magnitude of effect, a dose response  
139 relationship (i.e., evidence of a gradient), and cross-population and cross-study-design consistency. The  
140 certainty of evidence summary is shown in Table S5 (Supplementary Material pg. 6-8).

141

## 142 *Statistical Analysis*

### 143 Standardisation

144 Following data extraction, all papers were combined for statistical analysis. Given that some of the included  
145 studies used the same study population, we only included the most recent study with the longest follow-up  
146 in meta-analysis to avoid double counting. Meta-analyses were conducted when three or more independent  
147 studies reported an exposure-outcome pair (i.e., PM<sub>2.5</sub> and Dementia, PM<sub>10</sub> and Dementia). For each  
148 pollutant, adjusted and unadjusted effect estimates, and their 95% CI were harmonised to a prespecified  
149 exposure increment representing World Health Organization Standards (where no annual WHO guideline  
150 values exist, per the protocol, we calculated new concentration increments to approximately equal to 10%  
151 of the maximum concentrations encountered in any of the included studies) and entered in a Hartung and

152 Knapp random-effects model (23). Conversions were conducted using the approach utilized Chen and Hoek  
153 (2020), which has previously been used in various meta-analyses related to air pollution exposure (25). This  
154 process is detailed in Table S6 (Supplementary Material pg. 9). Conversions between ppb and  $\mu\text{g}/\text{m}^3$  were  
155 completed for  $\text{NO}_2$ ,  $\text{NO}_x$ , and  $\text{BC}/\text{PM}_{2.5}$  absorbance using the following conversions factors from the Health  
156 Effects Institute Special Report: for  $\text{NO}_2$   $\times 1.88$ , for  $\text{NO}_x$   $\times 1.55$ , for  $\text{BC}/\text{PM}_{2.5}$  absorbance, values recorded in  
157  $10^{-5}/\text{m}$  were converted to  $\mu\text{g}/\text{m}^3$  using  $(\times 1.1)/1.25$  (26). Conversions for  $\text{O}_3$  (annual) between ppb and  
158  $\mu\text{g}/\text{m}^3$  were completed using  $\times 1.96$  using the conversion factor from the UK Department for Environment,  
159 Food & Rural Affairs (27).

160

#### 161 Meta-Analysis

162 The Der Simonian and Laird (DSL) inverse-variance method, with the Hartung Knapp (HK) modification  
163 was used to weight the harmonised effects estimates for directness of comparison with other meta-analyses  
164 (28). The standard DSL method does not take into account the uncertainty in the estimation of the between-  
165 studies variance, which may be high, particularly when the number of studies included in the meta-analysis  
166 is low, and hence could result in an inflated false positive rate (29). Therefore, we applied the DSL method  
167 with the HK modification, which has been shown to result in a more adequate error rate (30). Based on  
168 reasoning reported in previous articles investigating pollutant similarities, and reflected in the included  
169 studies, we treated BC and  $\text{PM}_{2.5}$  absorbance as interchangeable in this meta-analysis (31, 32). When outcomes  
170 are rare (i.e., <10-20%), the hazard ratio (HR) and odds ratio (OR) can be approximated to a risk ratio (RR)  
171 (33). Given that dementia is a rare outcome, pooling the HR and OR together in one meta-analysis in this  
172 case was considered acceptable. The analysis of unadjusted effect estimates was designed to give insight  
173 into the effects of (residual) confounding factors.

174

#### 175 Quantifying Heterogeneity

176 Between-study inconsistency was assessed using the  $I^2$  statistic and the Cochran Q-test (Q). 95% prediction  
177 intervals (PIs), calculated from the between-studies variance (tau-squared) and describing a range within

178 which 95% of true effects are expected to lie, were reported as a more intuitive and interpretable description  
179 of the amount of variability in the meta-analysis (34) (35).

180

#### 181 Subgroup Analyses

182 The published protocol outlined that we would conduct subgroup analyses by geographical location, sex,  
183 age, ethnicity, exposure assessment method, outcome definition and ascertainment (23). Not enough papers  
184 ( $\geq 3$ ) recorded data on sex, age, or ethnicity to conduct subgroup analyses. Therefore, subgroup analyses  
185 were undertaken based on continent, exposure assessment method, and outcome definition and  
186 ascertainment. Additionally, subgroup analyses by study-diagnosed dementia subtype were completed.  
187 While not all subgroups had more than three studies, subgroup analyses with at least two subcategories with  
188 three or more studies are commented on in this manuscript, as we considered this to allow for a somewhat  
189 meaningful comparison within those groups.

190

#### 191 Sensitivity Analyses

192 We performed the following sensitivity analyses: 1) excluding studies which were determined to have  
193 ‘probably high risk of bias’; 2) excluding studies which rely on patient self-report of physician diagnosis  
194 for outcome assessment, where the potential of reporting error is higher compared to administrative medical  
195 (e.g., hospital) records; 3) using the Paule-Mandel (PM) with the HK modification in place of the DSL,  
196 given reports that the DSL is negatively biased in scenarios with small studies and with a rare binary  
197 outcome (36); 4) excluding the study with the largest weight (smallest SE) from each meta-analysis; 5)  
198 excluding studies with models which were adjusted for co-morbidities. There were two sensitivity analyses  
199 which were planned per the protocol but could not be run primarily due to the lack of studies, as detailed  
200 in Table S1 (Supplementary Material pg. 1).

201

#### 202 Visualising the shape of the Exposure Response Functions (ERF)

203 As outlined in our protocol, we planned to formally assess the shape of a pollutant's ERF when four or more  
204 independent studies reported an exposure-outcome pair across the exposure range, and those studies  
205 included effect estimates for at least three exposure categories, excluding the reference category. However,  
206 none of the exposure-outcome pairs met these pre-specified criteria. As a result, we did not formally  
207 evaluate ERF shapes, per our protocol, but instead visualised the data, adjusting the thresholds of the three  
208 exposure categories to include, rather than exclude, the reference category. Given the limited data, we  
209 deemed simple visualisation in this instance more appropriate, reserving methods like cubic splines for  
210 future analysis when more robust data becomes available.

211

#### 212 Assessment of Publication Bias

213 We visually examined publication bias with funnel plots and conducted the Egger's linear regression test  
214 to estimate the potential publication bias in cases of ten or more studies, per the protocol.

215

#### 216 Visualisation

217 The results of all meta-analyses are displayed using forest-plots. The standardisation of effect estimates  
218 was only possible for continuous variables.

219

220 All data-analysis was conducted in R version 4.2.2 using the “meta”, “dmetar”, “metafor”, “tidyverse”,  
221 “readxl”, “ggplot2”, “ggeasy”, “reshape2”, “DescTools”, “tm”, “stringr”, and “dplyr” packages (37-39).

222

#### 223 Role of Funder

224 The funders had no role in the study design or in the collection, analysis, interpretation of data, writing of  
225 the manuscript, or decision to submit the article for publication.

## Results

### *Primary Results*

Our initial search in October 2022 identified 8319 unique publications across eight databases. From this, 141 papers were full-text screened. 107 of those papers were excluded because they were either non-dementia related or met one of the other exclusion criteria. 34 papers published from the inception of the search engine until October 2022, met our inclusion criteria and were included in data extraction. Our secondary search conducted in October 2023 identified 1005 unique publications. 137 papers were identified for full-text screening and 17 were included in data extraction (Figure 1). Across both searches, in total, 51 papers were included.

Table 1 provides a summary of each paper. Most studies began at age 55, 60, or 65. The maximum reported age of a participant was 115 and the minimum reported was 37, though not all studies recorded their minimum and maximum age. Reported follow-up periods ranged from a minimum of 3 to a maximum of 22.7 years, though most studies reported follow-up either as a median or mean.

20 studies were conducted in Europe, 17 in North America, 12 in Asia, and 2 in Australia. 43 studies reported on dementia, 26 on AD, 17 on VaD, 1 on frontotemporal dementia (FTD), and 1 on mixed VaD/AD. Whether each paper was included in the meta-analysis was also documented in Table 1.

Forty studies reported on PM<sub>2.5</sub>, 28 on NO<sub>2</sub>, 17 on PM<sub>10</sub>, 12 on NO<sub>x</sub>, 10 on BC/PM<sub>2.5</sub> absorbance, 10 on annual O<sub>3</sub> (O<sub>3</sub> was separated between warm season and annual), and 6 on PM<sub>2.5-10</sub>. All studies except 3 were cohort studies. Of those 3, 1 was a case-control study (Wu et al. (2015)), and 2 combined cohort data with a nested case-control (Li et al. (2019) and Yuchi et al. (2020)).

Exposure assessment methods varied across studies and were divided into seven categories, based on an assessment of included papers and relevant literature on exposure assessment methods (40-42). The most popular method was Land Use Regression Models (LURs). Other methods included dispersion models, chemical transport models (CTMs), spatial interpolation techniques, statistical or machine-learning models (separate from LUR), monitoring station data, and hybrid models which combined a number of these approaches in one.

Five studies relied on some form of patient self-report of physician dementia diagnosis (at times validated by medical records), while the remaining 46 used a variety of methods including review of medical records, neuropsychological exams, and validated algorithms of health records to determine if an individual had a diagnosis of dementia, AD, VaD, or FTD (43-47). The reviews of medical records were either based on DSM (Diagnostic and Statistical Manual of Mental Disorders) coding or ICD (International Classification of Diseases) billing codes.

The RoB and Certainty of Evidence results are displayed in Tables S4 and S5. Only seven studies recorded a probably high RoB in one category, either for the use of some form of self-report of a physician diagnosis, a failure to report some analyses outlined in their methods, or the use of a non-peer reviewed categorization of pollutant exposure (qualifying under the category of “other bias”) (43-49).

1 *Meta-analysis*

2 Several of the 51 papers used the UK Biobank population or other large cohort populations (Medicare,  
3 Betula Cohort, GEMS, 3C Population, EHS: Taiwan, SALSA, WHIMS, NHIRD). While all papers were  
4 included for data extraction and the systematic review, per the protocol, only the most recent paper with  
5 the longest follow-up was selected for the meta-analysis. This led to the removal of ten papers from the  
6 meta-analysis solely for reasons of population overlap: Andersson et al. (2018), Chen et al. (2023), Ran et  
7 al. B (2021), Yuan et al. (2023), Ma et al. (2023), Chang et al. (2014), Raichlen et al. (2022), Cacciottolo  
8 et al. (2017), Li et al. (2022), Dimakakou et al. (2020)) (7, 10, 47, 50-56). A detailed explanation of the  
9 exclusion due to population overlap is provided in S18 (Supplementary Material pg. 53-54).

10

11 Following the removal of papers based on the same population, 41 remained for the meta-analysis. An  
12 additional paper (Xie et al. 2023) was removed from the meta-analysis due to a failure to record a CI in  
13 their presentation of the data bringing the final number of papers to 40 (58). Of those 40, 34 recorded  
14 pollutant-exposure pairs that reached the threshold of 3 or more papers and where therefore recorded in the  
15 meta-analysis results. The 34 papers included in the meta-analyses originated from 11 countries: 15  
16 originated in North America, 10 in Europe, 7 from Asia, and 2 from Australia

17

18 **PM<sub>2.5</sub>**

19 For PM<sub>2.5</sub>, 21 studies were included in the meta-analysis. The pooled effect estimate per 5 µg/m<sup>3</sup> was 1.08  
20 (95% CI: 1.02 to 1.14; Figure 2). The I<sup>2</sup> was 95% and Q was statistically significant (p<0.01). The PI was  
21 1.00 to 1.16.

22

23 Two studies in the pooled analysis had wide CIs. These values were rescaled to per 1 µg/m<sup>3</sup> and then  
24 manually checked. The wide CIs were determined to be the result of scaling up; the CIs were wide for

25 small increases in PM<sub>2.5</sub> exposure (per 0.88 µg/m<sup>3</sup> for Grande et al. (2020); per 1 µg/m<sup>3</sup> for Sullivan et al.  
26 (2021)) compared to the other included studies and scaling them up to per 5 µg/m<sup>3</sup> made them even wider.

27

## 28 **NO<sub>2</sub>**

29 For NO<sub>2</sub>, 16 studies were included. The pooled effect estimate per 10 µg/m<sup>3</sup> was 1.03 (95% CI: 1.01 to  
30 1.05; Figure 3). The I<sup>2</sup> was 84% and Q was statistically significant (p<0.01). The PI was 1.00 to 1.07.

31

## 32 **NO<sub>x</sub>**

33 For NO<sub>x</sub>, 5 studies were included. The pooled effect estimate per 10 µg/m<sup>3</sup> was 1.05 (95% CI: 0.97 to 1.13;  
34 Figure 4). The I<sup>2</sup> was 44%, Q was not statistically significant (p=0.13), and the PI was 0.92 to 1.19.

35

## 36 **BC/PM<sub>2.5</sub> Absorbance**

37 For BC/PM<sub>2.5</sub> Absorbance, 6 studies were included. The pooled effect estimate per 1 µg/m<sup>3</sup> was of 1.13 (95%  
38 CI: 1.01 to 1.27; Figure 5). I<sup>2</sup> was 97%, Q was statistically significant (p<0.01), and the PI was 0.72 to 1.78.

39

## 40 **PM<sub>10</sub>**

41 For PM<sub>10</sub>, 4 studies were included. The pooled effect estimate per 15 µg/m<sup>3</sup> was 1.52 (95% CI: 0.80 to 2.87;  
42 Figure 6). I<sup>2</sup> was 82%, Q was statistically significant (p<0.01), and the PI was 0.20 to 11.61.

43

## 44 **O<sub>3</sub> (annual)**

45 For annual O<sub>3</sub>, 4 studies were included. The pooled effect estimate per 45 µg/m<sup>3</sup> was 0.82 (95% CI: 0.35 to  
46 1.92; Figure S3A; Supplementary Material pg. 16). I<sup>2</sup> was 69%, Q was statistically significant (p=0.02),  
47 and the PI was 0.17 to 3.95.

48

49 No other pollutants reached the threshold of three studies for meta-analysis. Across these primary analyses,  
50  $I^2$  was 69% and above for PM<sub>2.5</sub>, PM<sub>10</sub>, O<sub>3</sub>, NO<sub>2</sub>, BC/PM<sub>2.5</sub> Absorbance and 44% for NO<sub>x</sub>, implying that the  
51 variability in effect size estimates may be due to sampling error. However,  $I^2$  may be less reliable with  
52 fewer included studies. The PIs for PM<sub>2.5</sub> and NO<sub>2</sub> did not dip below one, lending further support to the  
53 significance of those results.

54

### 55 *Subgroup Analysis*

#### 56 By Subtype of Dementia

57 When analysed by study-diagnosed subtype of dementia, the pooled effect estimate for PM<sub>2.5</sub> was 1.09  
58 (95% CI: 1.01 to 1.19) for AD, while for VaD it was 1.50 (95% CI: 0.69 to 3.26) per 5 µg/m<sup>3</sup> (Figure S4A;  
59 Supplementary Material pg. 17). Heterogeneity was high in both subgroups. While larger in magnitude, the  
60 effect estimate for VaD was not statistically different from that for AD (p=0.13).

61

62 For NO<sub>2</sub>, the effect estimate for AD was 1.03 (95% CI: 0.93 to 1.15) while for VaD it was 1.13 (95% CI:  
63 1.02 to 1.26) per 10 µg/m<sup>3</sup> (Figure S4B; Supplementary Material pg. 18). Though different in magnitude,  
64 the difference was not statistically significant between the subgroups (p=0.20).

65

66 For BC/ PM<sub>2.5</sub> Absorbance the effect estimate for VaD was 1.26 (95% CI: 0.95 to 1.66) and was higher than  
67 AD at 1.14 (95% CI: 0.85 to 1.52) per 1 µg/m<sup>3</sup> (Figure S4C; Supplementary Material pg. 19). The difference  
68 was not statistically significant (p=0.41).

69

70 For PM<sub>10</sub>, the effect estimate for VaD was 1.01 (95% CI: 0.74 to 1.37) and was lower than AD at 1.20 (95%  
71 CI: 0.60 to 2.40) per 15 µg/m<sup>3</sup> (Figure S4D; Supplementary Material pg. 20). The difference was not  
72 statistically significant (p=0.31).

73

74 By Continent

75 When analysed by continent, the effect estimate per 5  $\mu\text{g}/\text{m}^3$   $\text{PM}_{2.5}$  was 1.39 (95% CI: 0.80 to 2.44) for  
76 Europe, and 1.07 (1.01 to 1.13) for North America, with substantial heterogeneity in each of the continents.  
77 While the effect estimates differed across continents, the difference was not statistically significant ( $p=0.56$ ;  
78 Figure S5A; Supplementary Material pg. 21). For  $\text{NO}_2$ , the effect estimate for Europe was 1.09 (95% CI:  
79 0.96 to 1.23), while for North America it was 1.02 (95% CI: 1.01 to 1.04) per 10  $\mu\text{g}/\text{m}^3$  (Figure S5B;  
80 Supplementary Material pg. 22). Inconsistency was lower in the European subgroup, but heterogeneity was  
81 high in both analyses. This difference between subgroups was not statistically significant. There were two  
82 few studies to comment on  $\text{BC}/\text{PM}_{2.5}$  absorbance (Figure S5C; Supplementary Material pg. 23).

83

84 By Exposure Assessment Methods

85 When analysed by exposure assessment method, the difference between the pooled effect estimates for  
86  $\text{PM}_{2.5}$  was not statistically significant ( $p=0.47$ ; S6A; Supplementary Material pg. 24). Only the pooled effect  
87 estimates from studies which used LUR models or hybrid models, the subgroups with the largest number  
88 of studies, remained close to statistically significant. The difference between subgroups for  $\text{NO}_2$  was close  
89 to statistically significant ( $p=0.07$ ), though only the LUR subgroup remained statistically significant on its  
90 own, and it was the category with the largest number of studies (S6B; Supplementary Material pg. 25).

91

92 By Outcome Ascertainment Methods

93 For  $\text{PM}_{2.5}$  the pooled effect estimates differed significantly ( $p=0.02$ ) between outcome ascertainment  
94 methods (S7A; Supplementary Material pg. 26). Only the pooled effect estimates from studies which used  
95 medical records and a hybrid method, the subgroup with the largest number of studies, remained statistically  
96 significant. The other subgroups of studies (used physician diagnosis, Administered Exams) had larger,  
97 non-statistically significant pooled effect estimates but a much smaller number of studies. For  $\text{NO}_2$ , the

98 pooled effect estimates did not differ significantly between outcome ascertainment methods ( $p=0.51$ ) and  
99 none of the subgroups remained statistically significant (S7B; Supplementary Material pg. 27).

100

#### 101 Additional Subgroup Analyses

102 Due to the small number of studies that met the inclusion criteria for some pollutants, no further subgroup  
103 analyses could be conducted for those pollutants.

104

#### 105 *Sensitivity Analyses*

106 Additionally, while we outlined that we would conduct sensitivity analyses by study design, this was not  
107 feasible as the three studies that used some form of case-control study design only recorded values for AD  
108 and VaD (Yuchi et al. (2020) only used nested case-control for their AD results) (59-61).

109

#### 110 Paule-Mandel Estimator instead of Der-Simonian and Laird

111 Use of the PM estimator instead of the DSL yielded variable changes in the effect estimates but no changes  
112 to our conclusions. The effect estimate for dementia and  $PM_{2.5}$  was much higher (1.43 (95% CI: 1.05 to  
113 1.96)) (S8A; Supplementary Material pg. 28). For dementia and  $PM_{10}$ , it was slightly lower (1.47 (95% CI:  
114 0.77 to 2.79) (S8B; Supplementary Material pg. 29). For  $NO_2$ , it was marginally higher (1.05 (95% CI: 1.01  
115 to 1.09)) (S8C; Supplementary Material pg. 30), while for  $NO_x$ , the PM only slightly raised the pooled  
116 effect estimate (1.06 (95% CI: 0.96 to 1.16)) (S8D; Supplementary Material pg. 31). In all PM cases, the  
117 confidence intervals and PIs were wider than with the DSL, except for  $PM_{10}$  where both the confidence  
118 interval and the PI were narrower.

119

#### 120 Removal of Self-report

121 Of the studies included in the meta-analysis, both Wood et al. (2022) and Trevenen et al. (2022) relied on  
122 patient self-report of a physician diagnosis, and both were removed from the analysis (the others: Parra et

123 al. (2022), Mukadam et al. (2022) and Dimakakou et al. (2020) also relied on self-report, validated with the  
124 UK Biobank algorithm, but were not included in the meta-analysis due to an overlap in population and/or  
125 insufficient number of studies for their exposure-outcome pairs).

126

127 For PM<sub>2.5</sub>, the removal of these two papers altered the CI of the pooled effect estimate from 1.08 (95% CI:  
128 1.02 to 1.14) to 1.08 (95% CI: 1.02 to 1.15) (S9A; Supplementary Material pg. 32). Given that Wood et al.  
129 (2022) was 35.8% of the weight in the meta-analysis of PM<sub>10</sub> and dementia, the pooled effect estimate  
130 changed from 1.52 (95% CI: 0.80 to 2.87) to 1.83 (95% CI: 1.31 to 2.55) with the removal of that study  
131 (S9B; Supplementary Material pg. 33). For NO<sub>2</sub>, the removal of both papers changed the CI of the pooled  
132 effect from 1.03 (95% CI: 1.01 to 1.05) to 1.03 (95% CI: 1.01 to 1.06) (S9C; Supplementary Material pg.  
133 34). For BC/ PM<sub>2.5</sub> absorbance the removal of Trevenen et al. (2022) increased the pooled effect estimate from  
134 1.13 (95% CI: 1.01 to 1.27) to 1.15 (95% CI: 1.00 to 1.32), however it rendered it non-statistically  
135 significant (S9D; Supplementary Material pg. 35). The meta-analysis could not be conducted for annual O<sub>3</sub>  
136 because with the removal of Wood et al., it no longer met the criteria for analysis. Trevenen et al. (2022)  
137 and Wood et al. (2022) did not appear in other meta-analyses.

138

### 139 Excluding Studies with High Risk of Bias

140 All five studies which relied on some form of self-report of a physician diagnosis received a probably high  
141 risk of bias rating in the detection domain: Wood et al. (2022), Trevenen et al. (2022), Parra et al. (2022),  
142 Mukadam et al. (2022), and Dimakakou et al. (2020). In addition, two other studies received a rating of  
143 probably high risk of bias in one domain (He et al. (2022) and Ilango et al. (2020)). A sensitivity analysis  
144 was conducted, per the protocol, by excluding those that reported a high risk of bias in one domain (He et  
145 al. (2022), Parra et al. (2022), Mukadam et al. (2022) and Dimakakou et al. (2022) did not appear in any  
146 primary meta-analyses). Removing Ilango et al. (2020), Trevenen et al. (2022), and Wood et al. (2022)  
147 altered the pooled effect estimate for PM<sub>2.5</sub> from 1.08 (95% CI: 1.02 to 1.14) to 1.08 (95% CI: 1.01 to 1.15)

148 (Figure S10A; Supplementary Material pg. 36). For NO<sub>2</sub>, removing Ilango et al. (2020), Trevenen et al.  
149 (2022), and Wood et al. (2022) altered the upper end of the confidence interval (from 1.03 (95% CI:1.01 to  
150 1.05) to 1.03 (95% CI: 1.01 to 1.06) (S10B; Supplementary Material pg. 37). For PM<sub>10</sub>, removing Wood et  
151 al. (2022), increased the effect estimate from 1.52 (95% CI: 0.80 to 2.87) to 1.83 (95% CI: 1.31 to 2.55),  
152 though the number of included studies reduced from four to three (Figure S10C; Supplementary Material  
153 pg. 38). For BC/PM<sub>2.5</sub> Absorbance, removing Wood et al. (2022) increased the pooled effect estimate, however  
154 it rendered it non-statistically significant (from 1.13 (95% CI: 1.01 to 1.27) to 1.15 (95% CI: 1.00 to 1.32)  
155 (Figure S10D; Supplementary Material pg. 39). No studies were removed from the NO<sub>x</sub> meta-analysis.

156

#### 157 Excluding the Study with the Largest Weight

158 Excluding the studies with the largest weights had very little impact on the overall pooled effect estimates  
159 for PM<sub>2.5</sub>, NO<sub>2</sub>, NO<sub>x</sub>, and O<sub>3</sub> as demonstrated in the leave one out sensitivity analysis (Figure S11A,B,C, F;  
160 Supplementary Material pgs. 40-41). For BC/PM<sub>2.5</sub> absorbance, omitting Shi et al. (2023), which was 24% of  
161 the weight, lowered the confidence interval from 1.13 (95% CI: 1.01 to 1.27) to 1.06 (95% CI: 0.96 to 1.17)  
162 per 1 µg/m<sup>3</sup>. Generally, the BC/PM<sub>2.5</sub> absorbance meta-analysis was sensitive to omitting individual papers in  
163 the leave one out sensitivity analysis (Figure S11D; Supplementary Material pg. 42). Omitting Wood et al.  
164 (2022) in the PM<sub>10</sub> analysis, which was 35.8% of the weight, shifted the pooled effect estimate away from  
165 the null from 1.52 (95% CI: 0.80 to 2.87) to 1.83 (95% CI: 1.31 to 2.55) (Figure S11E; Supplementary  
166 Material pg.42). Removing Zhang, B. et al. (2023) from the O<sub>3</sub> analysis lowered the pooled effect estimate  
167 from 0.82 (95%CI: 0.35 to 1.92) to 0.76 (95% CI: 0.12 to 4.90) (Figure S11F; Supplementary Material pg.  
168 42).

169

#### 170 Excluding Studies which Adjusted for Comorbidities

171 Excluding the studies which adjusted for comorbidities in their models from the meta-analysis had small  
172 effects on the overall pooled effect estimates for PM<sub>2.5</sub>, NO<sub>2</sub>, NO<sub>x</sub> and PM<sub>10</sub> and did not change our

173 conclusions (Figure S12 A-C and E; Supplementary Material pg.43,44). For BC/PM<sub>2.5</sub> absorbance, the removal  
174 of Yuchi et al. (2020) and Zhang, Z. et al. (2023) raised the effect estimate from 1.13 (95% CI: 1.01 to 1.27)  
175 to 1.18 (95% CI:1.00 to 1.39), though it rendered it non-significant (Figure S12D; Supplementary Material  
176 pg. 44).

177

#### 178 Assessments of Publication Bias

179 Only NO<sub>2</sub> and PM<sub>2.5</sub> and dementia reached the threshold of ten studies for assessing small studies effects  
180 and using Egger's linear regression test. The studies for PM<sub>2.5</sub> are asymmetrical, with a large concentration  
181 of studies near the top with very low standard error (Figure S13A; Supplementary Material pg. 45). There  
182 are a few studies on the right-hand side of the graph with no studies to balance it out in the left-hand,  
183 suggesting there may be studies missing due to publication bias. The results of Egger's linear regression  
184 test (2.29, p=0.03) provide further evidence of funnel plot asymmetry. In a symmetrical funnel plot, studies  
185 with negative or non-significant effect sizes would likely appear on the left-hand side of the plot. The  
186 asymmetry indicates that fewer studies are reporting smaller or negative effects, which could either be due  
187 to publication bias or scarcity of such studies due to the presence of a true positive effect. The NO<sub>2</sub> studies  
188 are less asymmetrical and there is a suggestion of missing studies on the left-hand-side of the plot, again  
189 indicating potential publication bias (Figure S13B; Supplementary Material pg.46). The intercept of 1.19  
190 (p=0.10) from Egger's linear regression test does not differ significantly from zero. Again, these results may  
191 indicate a true positive effect.

192

#### 193 Categorical Exposure Response Functions (ERFs)

194 Five studies reported categorical effect estimates for PM<sub>10</sub> and AD, though for differing ranges of exposures  
195 (Figure S14A; Supplementary Material pg. 47). Two studies reported monotonic relationships (He et al.  
196 (2022), and Wu et al. (2015)), while the remaining three reported non-monotonic relationships. Four studies  
197 reported categorical effect estimates for NO<sub>2</sub> and AD, again for differing ranges of exposures. One study

198 reported a monotonic relationship (Chen et al. (2023)) and the other three reported non-monotonic  
199 relationship (Figure S14B; Supplementary Material pg. 47). Four studies reported categorical effect  
200 estimates for PM<sub>2.5</sub> and AD. Two reported monotonic relationships (Chen et al. (2023), and He et al. (2022))  
201 and two reported non-monotonic relationships (Figure S14C; Supplementary Material pg. 48). We  
202 determined that the data would not yield a more meaningful synthesis than the visual presentation in Figures  
203 S14A-C. No other pollutant-outcome pairs reached the threshold of four papers (excluding studies with  
204 population overlap in line with our overall meta-analysis conduct).

205 **Discussion**

206 This systematic review and meta-analysis found that there is a positive and statistically significant  
207 association between PM<sub>2.5</sub>, NO<sub>2</sub>, BC/PM<sub>2.5</sub> absorbance, and the incident dementia in disease-free individuals at  
208 baseline. There was no evidence for an association with incident dementia for NO<sub>x</sub>, PM<sub>10</sub> and annual O<sub>3</sub>  
209 based on a limited number of studies. These results are in line with findings of recent meta-analyses which  
210 largely focused on the associations between PM<sub>2.5</sub> and dementia. We also produce new evidence in support  
211 of significant associations between NO<sub>2</sub>, BC/PM<sub>2.5</sub> absorbance and dementia.

212  
213 Abolhasani et al. reported a significant association between PM<sub>2.5</sub> and dementia (1.03 (95% CI: 1.02 to  
214 1.05) per 1 µg/m<sup>3</sup>), a borderline statistically significant result with NO<sub>2</sub> (1.03 (95% CI: 1.00 to 1.07) per 10  
215 µg/m<sup>3</sup>) and some evidence of association with NO<sub>x</sub> (1.05 (95% CI: 0.99 to 1.13) per 10 µg/m<sup>3</sup>) (18).  
216 Similarly, Wilker et al. found evidence of an association between PM<sub>2.5</sub> and dementia (1.04 (95% CI: 0.99  
217 to 1.09) per 2 µg/m<sup>3</sup>), with higher associations when restricting the meta-analysis to the 7 studies that used  
218 active case ascertainment (1.42 (95% CI: 1.00 to 2.02) per 2 µg/m<sup>3</sup>), and a suggestive but non-significant  
219 association with NO<sub>2</sub> (1.02 (95% CI: 0.98 to 1.06) per 10 µg/m<sup>3</sup>) and NO<sub>x</sub> (1.05 (95% CI: 0.98 to 1.13) per  
220 10 µg/m<sup>3</sup>) (19). Notably, our meta-analysis included more studies than in both these meta-analyses. This  
221 review therefore strengthens the accumulating evidence for an association with PM<sub>2.5</sub>, adding 7 studies to  
222 the most recent meta-analysis by Wilker et al. (2023), and it builds on recent insufficient, but suggestive,  
223 evidence for associations with NO<sub>2</sub> and NO<sub>x</sub>, and provides new evidence for BC/PM<sub>2.5</sub> absorbance.

224  
225 For all pollutants except NO<sub>x</sub>, there was significant heterogeneity across the included studies. This likely  
226 reflects the range in the exposure assessment methods, the outcome ascertainment methods, the  
227 demographic characteristics of the study populations, the varying geographical and socioeconomic contexts  
228 of the studies, differences in the air pollution composition, and potentially underlying susceptibility factors.  
229 Large parts of this may be due to a lack of standard methods to assess exposure to air pollution in addition

230 to the limitations of standard methods such as their inability to capture personal mobility and associated  
231 exposures, the effects the clinical heterogeneity of dementia has on standardising case ascertainment and  
232 diagnosis, as well as potentially true differences in response. Subgroup analyses revealed that the pooled  
233 effect estimate PM<sub>2.5</sub>, BC/PM<sub>2.5</sub> absorbance, and NO<sub>2</sub> and VaD were higher compared to AD, although not  
234 statistically significant. Meta-analyses had previously been completed for PM<sub>2.5</sub> and AD/VaD and found  
235 similar results, but not for the range of pollutants analysed here (62).

236

237 The subgroup analyses by continent indicate that there are also potential differences in NO<sub>2</sub>, PM<sub>2.5</sub> and  
238 BC/PM<sub>2.5</sub> Absorbance effect estimates between continents, though none of the differences were statistically  
239 significant. For NO<sub>2</sub> in particular, the difference may speak to different NO<sub>2</sub> sources across continents, for  
240 example, the larger share of diesel vehicles in Europe versus in North America. Additional subgroup  
241 analyses by exposure assessment method indicate differences in effect estimates between exposure  
242 assessment methods for NO<sub>2</sub> and PM<sub>2.5</sub> with a close to statistically significant result (p=0.07 for NO<sub>2</sub>).  
243 Differences in pooled effect estimates were also noted between outcome ascertainment methods for both  
244 pollutants, with the difference reaching statistical significance in the case of PM<sub>2.5</sub>. These results indicate  
245 that additional research into exposure assessment methods prior to future meta-analyses may be necessary  
246 to assess whether certain exposure assessment perform better for capturing effects, and further attempts  
247 should be made to capture the true variance in the exposure of participants. Initial reports indicate that the  
248 performance of LUR, Bayesian maximum entropy (BME), or mixed models is affected by the temporal  
249 scale and degree of spatial heterogeneity. Therefore, the preferred model for a given environment depends  
250 on several contextual factors (63-65). The results on outcome ascertainment method may reflect the  
251 differences in the number of studies using each outcome ascertainment method (with those not reaching  
252 statistical significance having fewer studies). However, future work should focus on producing high quality  
253 reproducible outcome ascertainment methods that can be applied across different populations.

254

255 The results from the sensitivity analysis using the PM estimator were very similar to those in the primary  
256 analysis, except for PM<sub>2.5</sub>, where the point estimate of the effect size was higher and the confidence interval  
257 wider. It has previously been shown that in some circumstances the PM estimator can introduce substantial  
258 positive bias while the DSL method may be negatively biased, so the true effect may lie somewhere in  
259 between (66). The remaining sensitivity analyses supported this paper's primary findings for all pollutants  
260 except BC/PM<sub>2.5</sub> absorbance, as removing studies with patient self-report of physician diagnosis, studies with a  
261 rating of probably high risk of bias in one domain, studies with co-morbidities, and studies with the largest  
262 weight in the meta-analysis did not alter the conclusions. However, for BC/PM<sub>2.5</sub> absorbance, removing the  
263 studies with self-report of physician diagnosis, studies with a rating of probably high risk of bias in one  
264 domain, and those with co-morbidities made the results non-significant, indicating that those results should  
265 be viewed more critically. Finally, the leave one out sensitivity analysis showed rather robust pooled effect  
266 estimates for all pollutants, except for PM<sub>10</sub> and BC/PM<sub>2.5</sub> Absorbance, which was sensitive to the removal of  
267 individual studies. Future work when more data on the associations with BC/PM<sub>2.5</sub> Absorbance becomes  
268 available will be useful in elucidating that association.

269

### 270 *Findings in Context*

271 Epidemiological evidence plays a crucial role in building evidence for a causative relationship between air  
272 pollution and dementia. Our findings contribute to this case for a causal relationship, and complement  
273 evidence for the biological plausibility of the association (67). Several mechanisms have been proposed to  
274 explain how air pollution may cause dementia, though this remains an active area of research, and a  
275 comprehensive review of these mechanisms is beyond the scope of this paper. These mechanisms primarily  
276 involve neuroinflammation and oxidative stress, both of which play a well-established role in the  
277 pathogenesis of dementia (68-70). Air pollution is thought to trigger these processes through direct entry to  
278 the brain or via the same systemic effects that impact the risk of lung and cardiovascular diseases (71-73).

279

280 Neuropathological evidence of direct brain entry of air pollution exists across species in both animal and  
281 human studies. Dogs living in cities with high levels of air pollution were found to have metal accumulation  
282 in a gradient from olfactory mucosa to frontal cortex, suggesting the olfactory mucosa and nerve as a point  
283 of entry (73, 74). Blood-brain barrier disruptions, reactive astrocytosis, and neurofibrillary tangles are seen  
284 on electron microscopy of samples of the dogs' brains, matching the pathology seen in brains of humans  
285 with AD (74). Moreover, in rodent models, exposures to certain air pollution components were found to  
286 cause transcriptional activation of NF- $\kappa$ B, a key pro-inflammatory transcription factor, which may promote  
287 sustained production of neurotoxins such as reactive oxygen species through downstream activation of  
288 pathways including enzymes such as NADPH oxidases (72). Similar findings have been replicated in  
289 humans: on autopsy, compared to controls, children and young adults in a city with high air pollution have  
290 visible particulate matter in their olfactory bulb and increased expression of pro-inflammatory factors  
291 including cyclooxygenase-2, interleukin-1-beta, NF $\kappa$ B, and CD14 in their olfactory bulb, frontal cortex,  
292 substantia nigrae, and/or vagus nerves (75). Other cases saw evidence of tau hyperphosphorylation and  
293 amyloid-beta plaques, increased expression of NF $\kappa$ B, tumour necrosis factor (TNF), interferon (IFN) and  
294 toll-like receptors, as well as frontal disruption of the blood brain barrier (75, 76).

295

296 Beyond direct brain entry, systemic pathways may also contribute to the pathogenesis of dementia. Air  
297 pollutants can enter circulation from the lungs and travel to solid organs, initiating local and wide-  
298 spread inflammation (77-79). Ozone provides an example of how transcriptional modification may occur  
299 via indirect effects. There is evidence from mouse models that ozone causes a peripheral immune response  
300 after entry through the lungs, involving the upregulation of factors such as the High Mobility Group Box 1  
301 (HMGB1) protein (72, 80). This impairs the protective microglial response in the brain by reducing the  
302 expression of the triggering receptor expressed on myeloid cells 2 (TREM2), leading to increased  
303 accumulation of amyloid-beta plaques in a dose-dependent manner. Recently, research identified additional  
304 mechanisms, such as increased pulmonary production of lipopolysaccharide (LPS), which disrupts

305 microglial signalling and polarisation. This adds to the growing evidence that peripheral circulating factors  
306 contribute to microglial dysregulation (81, 82). Air pollution has also been found to directly induce protein  
307 misfolding, which may contribute to neurodegeneration through the endoplasmic reticulum stress response  
308 (83).

309  
310 With relevance to VaD, previous literature has demonstrated the strong association between air pollution and  
311 cardiovascular disease and stroke; this may represent a significant intermediate mediator of the effect of air  
312 pollution which may combine with direct effects (49, 84, 85). However, it is worth noting that alongside  
313 causative mechanisms which are thought to be unique to AD, cardiovascular disease is also associated with AD,  
314 and not just VaD (86, 87). Our findings therefore contribute to building the evidence-base for a causal link  
315 between air pollution and dementia, including AD and VaD, by providing epidemiological evidence supporting  
316 putative and biologically plausible mechanisms of air pollution's action on neurodegeneration (67).

317  
318 The effect estimates for AD were smaller than for VaD (except in the case of PM<sub>10</sub>), although the difference  
319 between these subgroups was non-significant. However, there are several important considerations to  
320 contemplate when interpreting this finding. Firstly, the subgroup analyses are still limited by the small  
321 number of included studies and generally lack statistical power. Secondly, most of the studies do not make  
322 a pathological diagnosis of either AD or VaD, but rather rely on non-validated, non-standardised methods  
323 of diagnosis which are not the clinicopathological diagnosis required to make an accurate subtype grouping  
324 (86, 88). Even if a pathological diagnosis is present, there is often poor correlation between pathological  
325 features on post mortem and a clinical diagnosis (21). Thirdly, there is increasing recognition that AD and  
326 VaD, although distinct clinical entities, often co-exist as a form of 'mixed dementia' and likely exist at  
327 different points on a single spectrum, and potentiate each other (21, 89-91). This means that there are  
328 significant barriers to being able to conclude this is a true effect. Notwithstanding these concerns over the  
329 boundaries and diagnosis of these disease entities, there may be several additional explanations of this  
330 result.

331  
332 There may be differences in the rate of case ascertainment between VaD and AD in the included studies which  
333 may have contributed, at least partially, to this difference. This is made more likely by the fact that VaD and AD  
334 may follow different clinical trajectories, which affect the point at which a diagnosis is sought, the rate at which  
335 symptoms appear, the method of diagnosis, and whether someone is diagnosed with a particular subtype of  
336 dementia (92, 93). Accordingly, we cannot assume that we have captured the true effect of these pollutants on  
337 the development of AD and VaD to the same degree (86, 87). Notably, there was substantial heterogeneity in  
338 the effect estimates with respect to AD, but generally less for VaD (except in the case of PM<sub>2.5</sub> where both VaD  
339 and AD had high heterogeneity). This suggests that there may be considerable heterogeneity in the clinical aspect  
340 of the disease, or the methodology of assessing exposure and outcome, or likely both. It is unclear why this  
341 heterogeneity seems more prominent for AD and whether additional data would paint a different picture; this is  
342 made even more surprising by the fact that VaD is generally more variable in its clinical presentation and has  
343 diagnostic criteria which has evolved more recently when compared to AD (94, 95). This difference seems to  
344 exist even within the context of a lack of overall standardisation of dementia subtype diagnosis. However, this  
345 subgroup analysis is informative to the extent that it is consistent with the fact that vascular risk factors and  
346 cardiovascular disease, including stroke, are often used to make a differential diagnosis of VaD (96), and  
347 therefore, may represent the recognised impact of air pollution on cardiovascular disease (85). We emphasise  
348 the need to develop consistent and standardised ways of dementia case ascertainment and more accurate methods  
349 for subtype grouping.

350  
351 We also investigated how the association of NO<sub>2</sub>, BC/PM<sub>2.5</sub> absorbance, and PM<sub>2.5</sub> with dementia varied  
352 by continent, although these differences were not statistically significant. This difference in the effect of  
353 could be due to the sources of air pollution. For example, while the transport sector is the largest source of  
354 NO<sub>2</sub> worldwide, more diesel than gasoline is burned in Europe compared to North America and a typical  
355 diesel car emits much more NO<sub>2</sub> than a gasoline-powered car (54, 55). These results replicate findings by  
356 Wilker et al. (2023), who also showed higher effect estimates in Europe when compared to North America

357 and noted that low exposure variance studies were based in Europe which may have led to the larger effect  
358 estimates in that region (19). Further investigation is necessary to understand the source of this difference.

359

### 360 *Limitations*

361 Firstly, there are limitations with respect to the studies included. A substantial limitation in air pollution  
362 and health research lies in the modelling of human exposure. Although assessment methods may effectively  
363 estimate air pollution exposure at the selected spatial resolution, they fail to capture individual/personal  
364 exposure variations due to individual movements throughout time and space and differences in dosage from  
365 varying inhalation rates by age, sex, race, nostril shape, and physical activity levels which all contribute to  
366 particle deposition in the body (75). The use of home address as a proxy for exposure assessment disregards  
367 time spent at work, other settings, or during commuting, which are exposure micro-environment that vary  
368 between individuals (97-99). This limitation may bias the results towards the null. Despite individual-level  
369 limitations, the pooled effect estimates derived from this analysis provide valuable insights into the  
370 population-level impacts and would allow the estimation of potential benefits of regulatory interventions.  
371 There is a need, however, to further refine exposure assessment methodologies and enhance their accuracy,  
372 reliability and representation of more complex real-world conditions.

373

374 The exposure assessment methods used, and the durations they are modelled for, additionally introduce  
375 uncertainty to the results presented in this study. The included papers used a variety of exposure assessment  
376 methods (including LUR, Dispersion models, CTM, etc.). Each model incorporates various uncertainties  
377 based on their set-up, input data, calculations, and underlying assumptions, and they accordingly have  
378 varying performances, including by air pollutant, and are fit for different purposes and geographical  
379 contexts. In addition to uncertainties stemming from the choice of exposure assessment methods, further  
380 uncertainty arises from the differences in time windows used in the included studies. Exposure windows  
381 ranged from a minimum of one year to a maximum of 22 years, with studies representing nearly every

382 duration within this range (44). Additionally, several studies also reported effect estimates for multiple  
383 exposure windows (e.g., 1, 3, and 5 years). These studies which provide such data include Shim et al.  
384 (2023), Li et al. (2019), and Zhang, B et al. (2023) (59, 100, 101). Per the protocol, we selected the model  
385 which related to the longest or most cumulative exposure window during the follow-up period with the  
386 assumption that this is the most relevant biological window. However, with the data we extracted, future  
387 research could assess how the different exposure windows within these papers compared to provide a more  
388 nuanced discussion of relevant exposure windows and their influence on the reported associations.  
389 Exploring the impact of these windows across different pollutants could provide insights into any varying  
390 temporal sensitivity of the associations. Since the confidence intervals in this meta-analysis do not account  
391 for uncertainty stemming from the different exposure assessment methods, they should be interpreted with  
392 caution, as they may underestimate the true extent of uncertainty present in models.

393

394 There are known and widespread limitations in the commonly used exposure assessment methods to  
395 examine the health effects of air pollutants. Specifically, although this paper looks at single pollutant  
396 models, air pollution exists as a complex mixture of multiple correlated pollutants, and the measurements  
397 of pollutants do not exist as mutually distinct categories. Although we may be tempted to conclude that  
398 some pollutants have stronger evidence linking them to dementia than others, an important consideration is  
399 the validity of the almost ubiquitous approach of treating exposure to each pollutant as occurring in isolation  
400 and being independent of other pollutants. This caution should frame our results. Exposure to air pollution  
401 occurs almost exclusively as a mixture, which makes interpretation of single pollutant models challenging.  
402 Reporting an association between a single pollutant and dementia may disregard the presence of interactions  
403 between pollutants, including permissive and synergistic effects (102-104). Categories of air pollutants are  
404 not mutually distinct: for example, major contributors and precursors to particulate matter include nitrogen  
405 oxides (105). This represents a major gap in our understanding and shortcoming of our methodology, which  
406 will be challenging to address: various approaches have been suggested but are challenging to implement

407 partly due to the correlated nature of exposures (102, 106-108). Per the protocol, if three or more studies  
408 reported on the same multi-pollutant models, we aimed to conduct an additional multi-pollutant assessment.  
409 The only multi-pollutant model for which there were sufficient studies to explore the relationship was PM<sub>2.5</sub>,  
410 controlling for NO<sub>2</sub> (Yu et al. (2023), Smargiassi et al. (2020), Yan et al. (2022)). We find no evidence of  
411 a significant relationship in this pooled analysis (1.02, 95% CI: 0.99 to 1.05) (S15A; Supplementary  
412 Material p.49). We recommend future research undertake this to investigate and further assess how the  
413 multi-pollutant models compare to the single-pollutant models that were included in this review to provide  
414 more robust insights into the relationship between these pollutants and the outcome of dementia.

415  
416 Furthermore, the papers with data on participants' ethnicity revealed that they were majority White. This is  
417 a limitation for the generalisability of results since there is evidence that marginalised groups have a higher  
418 exposure to air pollution which is not fully explained by socioeconomic status or geography (109, 110).  
419 This difference is made more significant by the evidence that reducing air pollution exposure has a higher  
420 benefit for reducing mortality for such marginalised groups (111, 112). Three studies reported effect  
421 estimates by ethnic group Mukadam et al. (2022), in their analysis for PM<sub>2.5</sub> and dementia, found a higher  
422 effect estimate for the Black population (1.38 (95% CI: 0.82 to 2.34) for those above 10 ug/m<sup>3</sup> compared  
423 to those below 10 ug/m<sup>3</sup>) compared to White (1.06 (95% CI: 1.00-1.12)) and South Asian populations (0.97  
424 (95% CI: 0.97 to 1.66)) (46). Younan et al. (2022), in their analysis of PM<sub>2.5</sub> and AD, similarly found a  
425 higher effect estimate for the Black population (2.11 (95% CI: 1.30 to 3.44) per 3.73 ug/m<sup>3</sup>) compared to  
426 the White population (1.07 (95% CI: 0.84 to 1.38)) (113). In contrast, Zhang, B. et al. (2023) in their  
427 analysis of PM<sub>2.5</sub> and dementia found a higher effect estimate for non-hispanic White (1.13 (95% CI: 1.03  
428 to 1.23) per 3.7 ug/m<sup>3</sup>) compared to non-hispanic Black (1.03 (95% CI: 0.91 to 1.17)), and Hispanic (1.04  
429 (95% CI: 0.95 to 1.14)) (100). While these studies lend some insights into the potential differing association  
430 between air pollution and dementia incidence across ethnicity, future work should ensure better and more

431 adequate representation across ethnicities. Additionally, most of the studies are based in high-income  
432 countries.

433  
434 The above factors together limit the generalisability of our results, as the particular social and structural  
435 context, such as level of education and socioeconomic status, as well as the prevalence of other  
436 comorbidities and life expectancy of a population, mediates the effect of air pollution on dementia. This  
437 has implications for addressing inequalities in burden of disease across countries but the problem of  
438 addressing health inequalities is preceded by the problem of defining them and the current dearth of studies  
439 into risk factors such as air pollution for dementia in low- and middle-income countries.

440  
441 Secondly, there are limitations with respect to the meta-analysis process. The conversion of hazard ratios  
442 to the same exposure increment assumed constant ambient temperature and pressure, which may introduce  
443 inaccuracies due to natural fluctuations in temperature and pressure throughout the year (114). However,  
444 given the lack of specific information from the included studies, these conversions were deemed the most  
445 appropriate approach and follow standard analytical practices. There are also further relevant study  
446 population characteristics future work should investigate, such as whether rural and urban settings are  
447 equally represented and if observed effects differ between those settings. This represents a potential source  
448 of confounding bias, as well as being relevant to health inequalities. Similarly, we have recorded whether  
449 studies adjusted for socioeconomic status, and at which level, but have not extracted the data to enable the  
450 investigation of effect estimates stratified by socioeconomic status.

451  
452 The adjusted meta-analysis included hazard ratios that were adjusted for different covariates and at varying  
453 levels (as described in Table 1). These ranged from clinical factors such as the presence of co-morbidities  
454 and body mass index, to lifestyle factors such as smoking and drinking, and socioeconomic factors. For  
455 example, some adjusted for individual participant income, while others adjusted for neighbourhood-level

456 socioeconomic position. This variation in adjustment methods limits the comparability of the effect  
457 estimates and could be investigated with meta-regression in future work. To partially address this issue, we  
458 also conducted preliminary analyses with the extracted unadjusted effect estimates when three or more were  
459 available. For PM<sub>2.5</sub> and dementia, the pooled effect estimate for the unadjusted values was 1.20 (95% CI:  
460 0.95 to 1.52) with eight papers (Figure S16A; Supplementary Material pg. 50). For NO<sub>2</sub>, the pooled effect  
461 estimate for the unadjusted values was 1.05 (95% CI: 0.93 to 1.18) with five papers (Figure S16B;  
462 Supplementary Material pg. 51). That the unadjusted effect estimates were higher for both NO<sub>2</sub> and PM<sub>2.5</sub>  
463 suggest that confounding factors were inflating the apparent association. Given that the number of studies  
464 in the unadjusted analysis is much smaller than the primary analysis, our interpretation remains very limited.  
465  
466 Finally, although our data extraction process adhered to gold-standard protocols with two independent  
467 reviewers (CR and YS), an internal review (HK) identified discrepancies, prompting a third independent  
468 verification of all extracted data. For instance, Oudin et al. (2018) was initially excluded because the  
469 original reviewers agreed the study, focusing on indoor stoves, did not meet the criteria for outdoor air  
470 pollution exposure (115). However, a subsequent review revealed the study included a model addressing  
471 outdoor air pollution, warranting its inclusion. Conversely, Tan et al. (2021) was initially included under  
472 the assumption that participants were disease-free at baseline, but further scrutiny found this could not be  
473 confirmed (116). These findings underscore the importance of additional verification steps beyond gold-  
474 standard practices. Requiring a third unaffiliated reviewer as an external check could enhance accuracy,  
475 although it added significant time and resource demands. Future opportunities may lie in leveraging  
476 artificial intelligence tools to streamline error-prone tasks, serving as an impartial reviewer, and potentially  
477 reducing the burden of manual oversight (117).

478

479 *Assessment of Confidence in the Body of Evidence*

480 We developed and published a protocol for this systematic review and meta-analysis a priori, and followed  
481 this precisely, lowering the risk of bias, and addressing factors that can decrease confidence in the body of  
482 evidence: “indirectness” and “imprecision” (23). The risk of bias of individual studies, as assessed by the  
483 OHAT tool, across the studies was at worst ‘probably low risk of bias’, with only seven having a ‘probably  
484 high risk of bias’ in one domain due to the use of some form of self-report of physician diagnosis (though  
485 at times verified through an algorithm) (Dimakakou et al. (2020), Parra et al. (2022), Trevenen et al. (2022),  
486 Mukadam et al. (2022), Wood et al. (2022)), failure to record certain information (Ilango et al. (2020)), and  
487 the use of a non-peer reviewed blog for determining the cut-offs for “good, moderate, and poor” air quality  
488 (He et al. (2022)). We conducted a RoB assessment for each individual paper to assess the data collection,  
489 analysis, and processing at the most granular level possible. This provided us with a fine-tuned assessment  
490 of the potential bias of each included paper. We did not categorically address potential bias at the journal  
491 level, as we could not find a validated method by which to do so.

492

493 In our Certainty of Evidence Assessment, we found the overall confidence in the body of evidence to be  
494 moderate (Table S5; Supplementary Material pg. 6-8). There was unexplained inconsistency and  
495 heterogeneity across results and evidence of publication bias, and these factors decreased our confidence.  
496 A range of heterogeneity is expected in meta-analyses, and has been a hallmark of air pollution  
497 epidemiology, and is not necessarily of concern, given that predefined eligibility criteria for the meta-  
498 analysis were sound and that the data were correctly extracted by two independent reviewers and verified  
499 by a third (118). However, the presence of cross-population and cross-study design consistency, and the  
500 evidence of a dose-response relationship in ERFs, increased our confidence in the body of evidence to  
501 moderate (Table S5; Supplementary Material pg. 6-8).

502

503 Per the protocol, a sensitivity analysis was conducted removing the seven studies that recorded a probably  
504 high risk of bias. The removal of such papers had only a negligible effect on the pooled effect estimates.

505 Grouping and running the analysis separately for studies with “Definitely low risk of bias” and “Probably  
506 low risk of bias” combined, versus those studies that included “Probably high risk of bias” also had a  
507 negligible effect (S17A,B; Supplementary Material pg. 52-53). Overall, the results of various sensitivity  
508 analyses led us to the same conclusions on the associations between pollutants and dementia.

509

510 As discussed, studies also maintained major differences in adjustment for confounders, outcome  
511 ascertainment and exposure assessment methods. Subgroup analyses were undertaken to assess those  
512 differences. Per the protocol, we aimed to increase comparability by selecting effect estimates from the  
513 papers which did not adjust for potential comorbidities, including cardiovascular disease. Finally, while our  
514 assessment of publication bias for those exposure-outcome pairs with more than ten papers indicate  
515 potential publication bias, given the presence of funnel plot asymmetry, these results may be true if there  
516 was a true positive effect estimate as well.

517

518 Taken together, the robustness of our protocol, the results of the certainty of evidence, risk of bias and  
519 various sensitivity and subgroup analyses, provide us with further confidence in the synthesised body of  
520 evidence and provide areas for further investigation.

521

### 522 *Implications*

523 Our results strengthen the evidence that outdoor air pollution is a risk factor for dementia. This suggests  
524 that efforts to reduce exposure to these pollutants in a population would help reduce the burden of dementia  
525 and that stricter limits for several pollutants are necessary with major contributors like the transport and  
526 industry sectors, signalling target areas for regulatory interventions. Reducing the air quality limits on  
527 PM<sub>2.5</sub>, NO<sub>2</sub>, to below current limits, and introducing limits on BC/PM<sub>2.5</sub> absorbance is likely to yield long-term  
528 health, social and economic benefits, relieving the healthcare system's burden by reducing dementia cases,  
529 income loss, and the significant toll on families and caregivers.

530

531 **Conclusion**

532 Our results strengthen and expand the evidence that  $PM_{2.5}$ ,  $NO_2$ , and  $BC/PM_{2.5}$  absorbance are risk factors for  
533 the onset of dementia including both study-diagnosed VaD and AD. This suggests that efforts to reduce  
534 exposure to these pollutants would help reduce the global burden of dementia. The extent of air pollution  
535 signals an urgent need for regional, national, and international policy interventions to combat air pollution  
536 equitably. Additionally, our meta-analysis reports on differences between effect estimates for AD and VaD  
537 and between continents. Interpretation of these differences remains limited by a limited data pool, and  
538 further research is necessary into the causal pathways and sources of air pollution, which may impact the  
539 development of dementia. Future work should aim to better represent low- and middle-income countries,  
540 include adequate diverse study populations, including diverse races and populations across varying degrees  
541 of urbanisation and socioeconomic status, to ensure global health inequalities are identified and addressed.  
542 This will ultimately inform a truly global and equitable approach to reduce the burden of dementia.

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## 891 **Declarations of Interest**

892 The authors declare that they have no known competing financial interests or personal  
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## 897 **Keywords**

898 Air pollution, dementia, environmental epidemiology, vascular dementia, Alzheimer's disease

## 899 **Figures Legend**

900 Figure 1: PRISMA Flow Diagram of Systematic Review Process

901  
902 Figure 2: PM<sub>2.5</sub> adjusted random-effects meta-analyses. Individual and summary random effects  
903 estimates for associations between PM<sub>2.5</sub> per 5 µg/m<sup>3</sup> and dementia. (HK) refers to the Hartung  
904 and Knapp random-effects model.

905 Figure 3: NO<sub>2</sub> adjusted random-effects meta-analyses. Individual and summary random effects  
906 estimates for associations between NO<sub>2</sub> per 10 µg/m<sup>3</sup> and dementia. (HK) refers to the Hartung  
907 and Knapp random-effects model.

908 Figure 4: NO<sub>x</sub> adjusted random-effects meta-analyses. Individual and summary random effects  
909 estimates for associations between NO<sub>x</sub> per 10 µg/m<sup>3</sup> and dementia. (HK) refers to the Hartung  
910 and Knapp random-effects model.

911  
912 Figure 5: BC/ PM<sub>2.5</sub> Absorbance adjusted random-effects meta-analyses. Individual and summary  
913 random effects estimates for associations between BC/ PM<sub>2.5</sub> Absorbance per 1 µg/m<sup>3</sup> and dementia.  
914 (HK) refers to the Hartung and Knapp random-effects model.

915 Figure 6: PM<sub>10</sub>, adjusted random-effects meta-analyses. Individual and summary random effects  
916 estimates for associations between PM<sub>10</sub> per 15 µg/m<sup>3</sup> and dementia. (HK) refers to the Hartung  
917 and Knapp random-effects model.

## 918 **Contributors**

919 CBR: Primary acquisition, analysis concept, analysis, and interpretation of data. Primary drafting  
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921 for important intellectual content. Administrative, technical, or material support.

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926 of long-term air pollution exposure on incidence of neurodegenerative diseases: a protocol for a  
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935 et al. (2022). Impact of long-term air pollution exposure on incidence of neurodegenerative  
936 diseases: a protocol for a systematic review and exposure-response meta-analysis. Environment  
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945 et al. (2022). Impact of long-term air pollution exposure on incidence of neurodegenerative

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971 term air pollution exposure on incidence of neurodegenerative diseases: a protocol for a  
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## 975 **Data Sharing Statement**

976 Data extracted for this systematic review and meta-analysis will be made available to others (as  
977 Table S9) with publication. The study protocol is already published (Khreis et al. (2022)) and  
978 accessible. The R code is available to be shared, should others seek to look at it and may be  
979 requested by contacting the corresponding author.